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ON A CLASS OF PERIMETER-TYPE DISTANCES OF PROBABILITY DISTRIBUTIONS

FERDINAND ÖSTERREICHER

The class I_{f_p} , $p \in (1, \infty]$, of *f*-divergences investigated in this paper generalizes an *f*-divergence introduced by the author in [9] and applied there and by Reschenhofer and Bomze [11] in different areas of hypotheses testing. The main result of the present paper ensures that, for every $p \in (1, \infty)$, the square root of the corresponding divergence defines a distance on the set of probability distributions. Thus it generalizes the respecting statement for p = 2 made in connection with Example 4 by Kafka, Österreicher and Vincze in [6].

From the former literature on the subject the maximal powers of f-divergences defining a distance are known for the subsequent classes. For the class of Hellinger-divergences given in terms of $f^{(s)}(u) = 1 + u - (u^s + u^{1-s})$, $s \in (0, 1)$, already Csiszár and Fischer [3] have shown that the maximal power is $\min(s, 1-s)$. For the following two classes the maximal power coincides with their parameter. The class given in terms of $f_{(\alpha)}(u) = |1 - u^{\alpha}|^{\frac{1}{\alpha}}$, $\alpha \in (0, 1]$, was investigated by Boekee [2]. The previous class and this one have the special case $s = \alpha = \frac{1}{2}$ in common. This famous case is attributed to Matusita [8]. The class given by $\varphi_{\alpha}(u) = |1 - u|^{\frac{1}{\alpha}} (1 + u)^{1 - \frac{1}{\alpha}}$, $\alpha \in (0, 1]$, and investigated in [6], Example 3, contains the wellknown special case $\alpha = \frac{1}{2}$ introduced by Vincze [13].

1. INTRODUCTION

Let (Ω, \mathcal{A}) be a nondegenerate measurable space (i.e. $|\mathcal{A}| > 2$ and hence $|\Omega| > 1$) and let $\mathcal{M}_1(\Omega, \mathcal{A})$ be the set of probability distributions on (Ω, \mathcal{A}) . Furthermore let \mathcal{F} be the set of convex functions $f : \mathbb{R}_+ \to \mathbb{R}$ which are continuous at 0. And let the function $f^* \in \mathcal{F}$ be defined by

$$f^*(u) = u \cdot f\left(\frac{1}{u}\right)$$
 for $u \in (0,\infty)$.

Remark 1. Owing to the continuity of f and f^* at 0 and by setting $0 \cdot f\left(\frac{0}{0}\right) = 0$ for all $f \in \mathcal{F}$ it holds

$$x \cdot f^*\left(\frac{y}{x}\right) = y \cdot f\left(\frac{x}{y}\right)$$
 for all $x, y \in \mathbb{R}_+$.

Definition (cf. Csiszár [4] and Ali and Silvey [1]). Let $Q_0, Q_1 \in \mathcal{M}_1(\Omega, \mathcal{A})$. Then

$$I_f(Q_1, Q_0) = \int f\left(\frac{q_1}{q_0}\right) \cdot q_0 \,\mathrm{d}\mu$$

is called f-divergence of Q_0 and Q_1 . (As usual, q_1 and q_0 denote the Radon-Nikodym-derivatives of Q_1 and Q_0 with respect to a dominating σ -finite measure μ .)

In the sequel we briefly restate those results from [6] which are basic for the statement and proof of the main result of this paper. For further informations on f-divergences we refer to the monograph [7] by Liese and Vajda and the paper [12] of Vajda and Österreicher.

Provided

(f1) f(1) = 0 and f is strictly convex at 1 and

(f2)
$$f^*(u) \equiv f(u)$$

it holds

 $\begin{array}{ll} (\mathrm{M1}) & I_f(Q_1,Q_0) \geq 0 \ \text{ with equality iff } Q_0 = Q_1 \ \forall \ Q_0,Q_1 \in \mathcal{M}_1(\Omega,\mathcal{A}) \,, \\ (\mathrm{M2}) & I_f(Q_1,Q_0) = I_f(Q_0,Q_1) & \forall \ Q_0,Q_1 \in \mathcal{M}_1(\Omega,\mathcal{A}) \end{array}$

respectively. If, in addition to (f1) and (f2), there exists an $\alpha \in (0, 1]$, such that

(f3,
$$\alpha$$
) the function $h(u) = \frac{(1-u^{\alpha})^{\frac{1}{\alpha}}}{f(u)}$, $u \in [0,1)$,

is (not neccessarily strictly) decreasing,

then, according to [6], Theorems 1 and 2, the power

$$\rho_{\alpha}(Q_0, Q_1) = [I_f(Q_1, Q_0)]^{\alpha}$$

of the f-divergence satisfies the triangle inequality

(M3) $\rho_{\alpha}(Q_0, Q_1) \leq \rho_{\alpha}(Q_0, Q_2) + \rho_{\alpha}(Q_2, Q_1) \quad \forall Q_0, Q_1, Q_2 \in \mathcal{M}_1(\Omega, \mathcal{A}).$

Remark 2. Note that by virtue of Jensen's inequality

$$\frac{f(u) + f^{*}(u)}{1 + u} = \frac{1}{1 + u} \cdot f(u) + \frac{u}{1 + u} \cdot f\left(\frac{1}{u}\right) \ge f(1) \,.$$

Therefore (f1) and (f2) imply f(u) > 0 for all $u \in \mathbb{R}_+ \setminus \{1\}$ and hence $f(0) \in (0, \infty)$. Moreover, it can be easily seen that, provided (f3, β) is satisfied for $\beta = \alpha \in (0, 1]$, it is also satisfied for every $\beta \in (0, 1]$.

The following Remark is a consequence of [6], Propositions 5 and 6.

Remark 3. Let (f1) and (f2) hold true and let $\alpha_0 \in (0, 1]$ be the maximal α for which (f3, α) is satisfied. Then the following statement concerning α_0 can be made. Let $k_0, k_1, c_0, c_1 \in (0, \infty)$ be such that

$$\begin{aligned} f(0) \cdot (1+u) - f(u) &\sim c_0 \cdot u^{k_0} & \text{for } u \downarrow 0 \quad \text{and} \\ f(u) &\sim c_1 \cdot |u-1|^{k_1} & \text{for } u \uparrow 1 \end{aligned}$$

then $k_0 \leq 1$, $k_1 \geq 1$ and $\alpha_0 \leq \min\left(k_0, \frac{1}{k_1}\right) \leq 1$.

2. THE MAIN RESULT

First we are going to show that f-divergences can be defined in terms of the following class of functions

$$f_p(u) = \begin{cases} (1+u^p)^{\frac{1}{p}} - 2^{\frac{1}{p}-1} \cdot (1+u) & \text{for } p \in (1,\infty) \\ \frac{|u-1|}{2} & \text{for } p = \infty \end{cases}, \quad u \in \mathbb{R}_+$$

which satisfies $\lim_{p\to\infty} f_p(u) = f_{\infty}(u)$.

Lemma 1. $f_p \in \mathcal{F}$ and satisfies (f1) and (f2) for all $p \in (1, \infty]$.

Proof. Since this assertion is obvious for the case $p = \infty$, let us assume $p \in (1,\infty)$ from now on. For this case

$$\lim_{u \downarrow 0} f_p(u) = f_p(0) = 1 - 2^{\frac{1}{p} - 1} \in (0, \infty), \quad f_p(1) = 0,$$

(1)

$$f'_p(u) = (1+u^p)^{\frac{1}{p}-1} \cdot u^{p-1} - 2^{\frac{1}{p}-1}$$
 and hence
 $f'_p(1) = 0$ and

(2)

$$f_p''(u) = (p-1) \cdot (1+u^p)^{\frac{1}{p}-2} \cdot u^{p-2} > 0 \quad \forall \ u \in (0,\infty) \quad \text{and hence}$$
$$f_p''(1) = (p-1) \cdot 2^{\frac{1}{p}-2}.$$

Therefore f_p is an element of \mathcal{F} satisfying (f1). The validity of $f_p^*(u) \equiv f_p(u)$ is obvious.

Remark 3 provides an upper bound for the subset of those $\alpha \in (0, 1]$, for which $(f_{3,\alpha})$ may hold.

Remark 4. Owing to

$$\begin{aligned} f_p(0) \cdot (1+u) - f_p(u) &= 1 + u - (1+u^p)^{\frac{1}{p}} \sim u & \text{for } u \downarrow 0 \,, \\ f_p(u) &\sim (p-1) \cdot 2^{\frac{1}{p}-3} \cdot (u-1)^2 & \text{for } u \uparrow 1 \,, \end{aligned}$$

(the latter being a consequence of (1) and (2)), the maximal $\alpha \in (0, 1]$ satisfying $(f3, \alpha)$ – if there is any – must be $\alpha_0 \leq \frac{1}{2}$.

Interpretation of the f-divergences under consideration. Let

$$R(Q_0, Q_1) = co\{(Q_0(A), Q_1(A^c)), A \in \mathcal{A}\}\$$

be the risk set of the testing problem $(Q_0, Q_1) \in \mathcal{M}_1(\Omega, \mathcal{A})^2$ (whereby "co" means "the convex hull of"). Then the corresponding f-divergence

$$I_{f_p}(Q_1, Q_0) = \begin{cases} \int (q_1^p + q_0^p)^{\frac{1}{p}} d\mu - 2^{\frac{1}{p}} & \text{for } p \in (1, \infty) \\ \frac{1}{2} \int |q_1 - q_0| d\mu & \text{for } p = \infty \end{cases}$$

can be interpreted as the difference or the arc lengths of the lower boundary of the risk set and the diagonal

$$D = \{(x, y) \in [0, 1]^2 : x + y = 1\},\$$

both measured in terms of the l_p -norm in \mathbb{R}^2 . We denote the arc length in question by l_p -arc length since it coincides for p = 2 with the ordinary arc length. For further reading on the geometric point of view we refer to Feldman and Österreicher [5] and the entry [10] of the author.

For the limiting case $p = \infty$ the corresponding f-divergence $I_{f_{\infty}}(Q_1, Q_0)$ is half of the well-known variation distance. For p = 2 it has been shown in [6] that the square root of the corresponding f-divergence $I_{f_2}(Q_1, Q_0)$ is also a distance. In the sequel we are going to show the following generalization of the latter which may be conjectured from Remark 4.

Theorem. For every $p \in (1, \infty)$ the square root of the *f*-divergence $I_{f_p}(Q_1, Q_0)$ defines a distance on $\mathcal{M}_1(\Omega, \mathcal{A})$.

By virtue of Lemma 1 and [6], Theorems 1 and 2, the proof is reduced to that of the following Lemma.

Lemma 2. Let $p \in (1, \infty)$. Then the function

$$h_p(u) = rac{(\sqrt{u}-1)^2}{f_p(u)}, \quad u \in [0,1)\,,$$

is (strictly) decreasing.

Proof. Because of

$$h_p'(u) = \left(\frac{1}{\sqrt{u}} - 1\right) \cdot \frac{1}{f_p^2(u)} \cdot \phi_p(u)$$

with

$$\begin{split} \phi_p(u) &= -\left[f_p(u) + \left(\sqrt{u} - u\right) \cdot f_p'(u)\right] \\ &= 2^{\frac{1}{p}-1}(1+u^{\frac{1}{2}}) - (1+u^p)^{\frac{1}{p}-1} \cdot (1+u^{p-\frac{1}{2}}) \end{split}$$

it suffices to show $\phi_p(u) < 0$ for all $u \in (0, 1)$. Owing to $\phi_p(1) = 0$ it suffices to show that the functions ψ_p defined by $\psi_p(u) = \sqrt{u} \cdot \phi'_p(u)$ satisfy

$$\psi_p(u) = 2^{\frac{1}{p}-2} - (1+u^p)^{\frac{1}{p}-2} \cdot u^{p-1} \cdot \left[(p-1) \cdot (1-\sqrt{u}) + \frac{1+u^p}{2} \right] > 0$$

for all $u \in (0, 1)$. Because of $\psi_p(1) = 0$ this, however, follows from

$$\psi'_p(u) = -(p-1) \cdot (p-\frac{1}{2}) \cdot (1+u^p)^{\frac{1}{p}-3} \cdot u^{p-2} \cdot (1-\sqrt{u}) \cdot (1-u^p) < 0,$$

which is obvious.

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REFERENCES

- S. M. Ali and S. D. Silvey: A general class of coefficients of divergence of one distribution from another. J. Roy. Statist. Soc. Ser. B 28 (1966), 131-142.
- [2] D.E. Boekee: A generalization of the Fisher information measure. Delft University Press, Delft 1977.
- [3] I. Csiszár and J. Fischer: Informationsentfernungen im Raum der Wahrscheinlichkeitsverteilungen. Magyar Tud. Akad. Mat. Kutató Int. Kösl. 7 (1962), 159–180.
- [4] I. Csiszár: Eine informationstheoretische Ungleichung und ihre Anwendung auf den Beweis der Ergodizität von Markoffschen Ketten. Publ. Math. Inst. Hungar. Acad. Sci. 8 (1963), 85-107.
- [5] D. Feldman and F. Osterreicher: Divergenzen von Wahrscheinlichkeitsverteilungen integralgeometrisch betrachtet. Acta Math. Acad. Sci. Hungar. 37 (1981), 4, 329–337.
- [6] P. Kafka, F. Osterreicher and I. Vincze: On powers of f-divergences defining a distance. Studia Sci. Math. Hungar. 26 (1991), 415-422.
- [7] F. Liese and I. Vajda: Convex Statistical Distances. Teubner-Texte zur Mathematik, Band 95, Leipzig 1987.
- [8] K. Matusita: Decision rules based on the distance for problems of fit, two samples and estimation. Ann. Math. Stat. 26 (1955), 631-640.
- [9] F. Osterreicher: The construction of least favourable distributions is traceable to a minimal perimeter problem. Studia Sci. Math. Hungar. 17 (1982), 341-351.
- [10] F. Osterreicher: The risk set of a testing problem A vivid statistical tool. In: Trans. of the Eleventh Prague Conference, Academia, Prague 1992, Vol. A, pp. 175–188.
- [11] E. Reschenhofer and I. M. Bomze: Length tests for goodness of fit. Biometrika 78 (1991), 207-216.
- [12] I. Vajda and F. Osterreicher: Statistical information and discrimination. IEEE Trans. Inform. Theory 39 (1993), 3, 1036-1039.
- [13] I. Vincze: On the concept and measure of information contained in an observation. In: Contributions to Probability (J. Gani and V.F. Rohatgi, eds.), Academic Press, New York 1981, pp. 207-214.

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