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A MARTINGALE CENTRAL LIMIT THEOREM
Petr LACHOUT

Abstract: The paper presents a martingale central limit theorem which connects the well-known result by McLeish (1974) with that one by Hall and Heyde (1980) and continues the research starting in [2].

Key words and phrases: A zero-mean martingale array, the central limit theorem, a uniform integrability.

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Let us formulate main results.

Theorem: Let $(S_{nk}, A_{nk}, k=1, \dots, k_n, n \in \mathbb{N})$ be a zero-mean martingale array with differences X_{nk} . Suppose that

- (i) $E \max \{ |X_{nk}| \mid k=1, \dots, k_n \} \rightarrow 0$
- (ii) $U_n = \sum_{k=1}^{k_n} X_{nk}^2 \xrightarrow{d} \eta^2$, where η^2 is an a.s. finite random variable,

(iii) $\lim_{k \rightarrow +\infty} \limsup_{n \rightarrow +\infty} E[\exp(-tU_n) - E[\exp(-tU_n)/A_{nk}]] = 0$
for every positive number t .

Then $S_{nk_n} \xrightarrow{d} S$, where the r.v.S has the characteristic function $E \exp(-\frac{1}{2} t^2 \eta^2)$.

Proof: The proof has the same framework as the proof of

1) \xrightarrow{d} means the usual convergence in distribution.

the theorem (2.3) in [3] and as the proof of the theorem 3.2 in [1], chapter 3, p. 58.

Put $M_n = \max \{|X_{nk}| \mid k=1, \dots, k_n\}$ and fix a real number t and positive number ε . According to (iii) there are a natural number j and a real number D such that

- (1) $P(\eta^2 \geq D) < \varepsilon$,
 (2) $\limsup_{n \rightarrow +\infty} E|\exp(-\frac{t^2}{2} U_n) - E[\exp(-\frac{t^2}{2} U_n)/A_{nj}]| < \varepsilon \exp(-\frac{t^2}{2} D)$.

Define the following transformation

$$Y_{nk} = X_{nk} I\left(\sum_{s=1}^{k_n-1} X_{ns}^2 \leq D\right),$$

$$J_n = \begin{cases} \max\{k \mid Y_{nk} \neq 0\} & \text{if there is a natural number } k \text{ such} \\ & \text{that } Y_{nk} \neq 0, \\ J & \text{if } Y_{nk} = 0 \text{ for every } k=1, \dots, k_n. \end{cases}$$

$(Y_{nk}, A_{nk}, k=1, \dots, k_n, n \in \mathbb{N})$ is obviously an array of martingale differences.

$$\text{Denote } T_{nk} = \prod_{s=1}^{k_n} (1 + itY_{ns}), \quad T_n = T_{nk_n},$$

$$W_n = \sum_{s=3}^{+\infty} \frac{(-it)^s}{s} \prod_{r=1}^{k_n} Y_{nr}^s,$$

$$B_n = [M_n \leq \frac{1}{2|t|}] \cdot F_n = [U_n \leq D], \quad C_n = B_n \cap F_n.$$

Now we can calculate

$$|W_n| \leq t^2 \sum_{s=3}^{+\infty} (|t|M_n)^{s-2} (Y_{nJ_n}^2 + \sum_{k=1}^{J_n-1} Y_{nk}^2) \leq t^2(M_n^2 + D) \sum_{s=3}^{+\infty} |t|^{s-2} s^{-2}.$$

Hence by (i)

- (3) $W_n I(B_n)$ are uniformly bounded r.v.'s and $W_n I(B_n) \xrightarrow{d} 0$.

We may derive an inequality for T_{nk}

$$|T_{nk}| \leq (1 + |t| |Y_{nJ_n}|) \prod_{s=1}^{J_n-1} (1 + t^2 Y_{nk}^2)^{\frac{1}{2}}$$

- (4) $|T_{nk}| \leq (1 + |t|M_n) \exp(\frac{1}{2} t^2 D)$.

We shall use the following property.

Lemma: Let f_n be complex functions which are A_{n_j} -measurable and uniformly bounded. Then $E(T_n - 1)f_n \rightarrow 0$.

Proof: $E T_n f_n = E \{ T_{n_j} f_n E [\prod_{k=j+1}^{n_j} (1 + it Y_{nk}) / A_{n_j}] \} = E T_{n_j} f_n$.
Then $E(T_n - 1)f_n \rightarrow 0$ since $T_{n_j} \xrightarrow{d} 1$. \square

Notice that

$$\begin{aligned} & E [T_n \exp(-\frac{t^2}{2} U_n) I(F_n)] - E \exp(-\frac{t^2}{2} U_n) = \\ & = E [T_n (\exp(-\frac{t^2}{2} U_n) - E[\exp(-\frac{t^2}{2} U_n) / A_{n_j}])] + \\ & + E \{ (T_n - 1) E[\exp(-\frac{t^2}{2} U_n) / A_{n_j}] \} - E [T_n \exp(-\frac{t^2}{2} U_n) I(U_n > D)]. \end{aligned}$$

Using (1), (2), (4) and the previous lemma we obtain

$$\begin{aligned} (5) \quad \limsup_{n \rightarrow +\infty} E | T_n \exp(-\frac{t^2}{2} U_n) I(F_n) - E \exp(-\frac{t^2}{2} U_n) | & \leq \\ & \leq 2\varepsilon + 2 |t| \exp(\frac{t^2}{2} D) \limsup_{n \rightarrow +\infty} E M_n = 2\varepsilon. \end{aligned}$$

Now we may write

$$\begin{aligned} & E \exp(it \sum_{k=1}^{n_j} \lambda_{nk}) - E \exp(-\frac{t^2}{2} \eta^2) = E [\exp(it \sum_{k=1}^{n_j} \lambda_{nk}) (1 - I(C_n))] + \\ & + E \{ \exp(it \sum_{k=1}^{n_j} \lambda_{nk}) - T_n \exp(-\frac{t^2}{2} U_n + W_n) \} I(C_n) + \\ & + E [T_n \exp(-\frac{t^2}{2} U_n) (\exp W_n - 1) I(C_n)] + E [T_n \exp(-\frac{t^2}{2} U_n) (I(C_n) - I(F_n))] + \\ & + \{ E [T_n \exp(-\frac{t^2}{2} U_n) I(F_n)] - E \exp(-\frac{t^2}{2} U_n) \} + \\ & + \{ E \exp(-\frac{t^2}{2} U_n) - E \exp(-\frac{t^2}{2} \eta^2) \}. \end{aligned}$$

Noting that the second term of the right hand side of the equality is vanishing, we can see

$$\begin{aligned} & | E \exp(it \sum_{k=1}^{n_j} \lambda_{nk}) - E \exp(-\frac{t^2}{2} \eta^2) | \leq P \cdot M_n > \frac{1}{2|t|} + P(U_n > D) + \\ & + E [|T_n| |\exp W_n - 1| I(B_n)] + E |T_n| I(M_n > \frac{1}{2|t|}) + \end{aligned}$$

$$+ |E[T_n \exp(-\frac{t^2}{2} U_n) I(F_n)] - E \exp(-\frac{t^2}{2} U_n)| +$$

$$+ |E \exp(-\frac{1}{2} t^2 U_n) - E \exp(-\frac{1}{2} t^2 \eta^2)|.$$

Using (i),(ii),(1),(3),(4) and (5) we obtain that

$$\limsup_{n \rightarrow +\infty} |E \exp(it \sum_{k=1}^{k_n} X_{nk}) - E \exp(-\frac{1}{2} t^2 \eta^2)| \leq 3\epsilon.$$

Now, it is clear that $S_{nk_n} \xrightarrow{d} S$, where the r.v. S has the characteristic function $E \exp(-\frac{1}{2} t^2 \eta^2)$. $\square \square$

Finally, let us remark that each of the following conditions implies the condition (iii).

(6) For every positive numbers ϵ, t there are a natural number j and functions f_n that are A_{nj} -measurable, $n \in N$, such that

$$\limsup_{n \rightarrow +\infty} E |\exp(-tU_n) - f_n| < \epsilon.$$

(7) Let ϵ be a positive number and $B_n \in \mathcal{G}(U_n)$, $n \in N$. Then there are a natural number j and sets $C_n \in A_{nj}$, $n \in N$, such that $P(B_n \Delta C_n) < \epsilon$ for any $n \in N$.

(8) η^2 is a nonnegative constant a.s.

(9) The martingale array is defined on a common probability space, $U_n \xrightarrow{D} \eta^2$ and the \mathcal{G} -fields A_{nk} are nested (i.e. $A_{nk} \subset A_{n+1,k}$ for $k=1, \dots, k_n, n \in N$).

Note that (8) is the assumption (c) of the theorem (2.3) in [3] and (9) are the assumptions (3.19) and (3.21) of the theorem 3.2 in [1].

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